# Independent Component Analysis

#### Uri Shaham

March 4, 2024

# 1 Whitening

Let  $X \in \mathbb{R}^m$  be a zero-mean random vector. Whitening linearly transforms X into  $\tilde{X}$ , so that the coordinates of  $\tilde{X}$  are uncorrelated and have unit variance, i.e.,  $\mathbb{E}\left[\tilde{X}\tilde{X}^T\right] = I$ . Let  $\mathbb{E}\left[XX^T\right] = V\Lambda V$  be the eigendecomposition of the covariance, so that  $V^T X$  is the projection of X onto its principal directions, as in PCA. The whitening transform is given by  $\tilde{X} = V\Lambda^{-\frac{1}{2}}V^T X$  (i.e., each principal component is scaled to have unit variance). Then

$$\mathbb{E}\left[\tilde{X}\tilde{X}^{T}\right] = V\Lambda^{-\frac{1}{2}}V^{T}\mathbb{E}\left[XX^{T}\right]V\Lambda^{-\frac{1}{2}}V^{T}$$
$$= V\Lambda^{-\frac{1}{2}}V^{T}V\Lambda V^{T}V\Lambda^{-\frac{1}{2}}V^{T}$$
$$= I$$



Figure 1: Example of whitening. Figure taken from https://www.cs.cmu.edu/~bapoczos/Classes/ML10715\_2015Fall/slides/ICA.pdf

**Remark 1.1.** The above procedure, with the rotation back (i.e., the leftmost multiplication by V) is sometimes called ZCA whitening. People often refer to whitening transform without the rotation back, i.e.,  $\tilde{X} = \Lambda^{-\frac{1}{2}} V^T X$  (known as PCA whitening). You will show in homework that if  $X_n = U \Sigma V^T$  is a  $n \times d$  data matrix, PCA whitening  $\Lambda^{-\frac{1}{2}} V^T X_n^T$  simply returns  $U^T$ .

### 2 Independent Component Analysis

Let  $S = (S_1, \ldots, S_n)^T$  be a vector of latent independent random variables (i.e.,  $\Pr(S) = \Pr(S_1, \ldots, S_n) = \prod_i \Pr(S_i)$ ), with zero mean and identity covariance. We observe *n* linear combinations of the latent random variables, given by X = AS, where  $A \in \mathbb{R}^{n \times n}$  is unknown. Our goal is to recover *S*, by computing  $W = A^{-1}$ .



Figure 2: Difference between PCA and ICA. Figure taken from https://hastie.su.domains/Papers/ icatalk.pdf

Suppose that  $S_1, \ldots, S_n$  are all standard Gaussian. Note that for every matrix B, BS is also Gaussian. Then for any  $n \times n$  rotation matrix R (i.e.,  $RR^T = I$ ) we have

$$\Pr(RS) = \frac{1}{2\pi} \exp\left(-\frac{S^T R^T RS}{(2R^T R)^{-1}}\right) = \frac{1}{2\pi} \exp\left(-\frac{S^T S}{2I}\right),$$

which means that s cannot be recovered (or put another way, A is not identifiable). Hence from now on we assume all variables are non-Gaussian.

### **3** Nongaussianity

Lyapunov's version of the central limit theorem asserts that sum of independent (not necessarily identically distributed) random variables converges in distribution to normal. Thus, intuitively, a  $X_j$ , which is the dot product between the *j*'th row of *A* and *S* is "more Gaussian" any of the  $S_i$ 's.

We want to recover one of the latent factors  $S_i$ , via  $Y := w^T X = (w^T A)S$ , which is a linear combination of the latent factors as well. Hence, to recover one of the components, we wish to find w

which maximizes the nonGaussianity of  $w^T X$ . A popular measure for nonGaussianity is negentropy, described next.

#### 3.1 Negentropy

**Definition 3.1.** The differential entropy of a random variable Y with density f is  $h(Y) := \int f(y) \log f(y) dy$ 

Fact 3.2. A Gaussian random variable has the largest entropy among all random variables with equal variance.

**Definition 3.3** (Negentropy). The Negentropy of a random variable Y is defined as  $J(Y) := h(Y_{Gauss}) - h(Y)$ , where  $h(Y_{Gauss}) = \frac{1}{2} \log (2\pi e\sigma)$  is the entropy of a Gaussian random variable with the same variance as Y.

Computing h(Y) is hard, as it requires a nonparametric estimation of the density f(Y). Hence, one typically use approximations for it. Specifically, negentropy is typically estimated by a non-quadratic function G (e.g.,  $G_1(y) = y^4$ ,  $G_2(y) = -\exp(y^2)$ , as

$$J(Y) \propto J(Y) := \left(\mathbb{E}[G(Y)] - \mathbb{E}[G(Z)]\right)^2,$$

where Z is a standard Gaussian random variable.

## 4 Solving ICA

We will aim to find an approximation Y of S. Since independent components are uncorrelated, we can restrict our search to matrices  $Y_n$  which are orthogonal, hence whitening can use as a starting point. Hence prior to the optimization, we preprocess the data matrix  $X_n$  by subtracting the mean from each column, following by whitening.

The minimization problem can be solved using standard methods, e.g., Newton's method

$$w^{(t+1)} = w^{(t)} - \left(\nabla^2 \tilde{J}(\tilde{X}_n w^{(t)})\right)^{-1} \nabla \tilde{J}(\tilde{X}_n w^{(t)}),$$

where expectations are replaced by sample means. For the first combination w, the requirement unit variance  $\operatorname{Var}\left(w^T \tilde{X}\right) = 1$ , together with the fact that  $\tilde{X}$  is whitened, is equivalent to requiring that w is a unit vector. This can be implemented by rescaling  $w_t$  after each iteration of the optimization procedure. For subsequent combination, we want each vector w to live in the orthogonal complement of the w's found so far, which we can achieve by applying Gram-Shmidt:

$$w_k \leftarrow w_k - \sum_{i=1}^{k-1} w_k^T w_i w_i.$$

## **Further Reading**

A good ICA tutorial is https://www.cs.jhu.edu/~ayuille/courses/Stat161-261-Spring14/HyvO00-icatut.pdf.